

CHAPTER 113

Noise and contrast discrimination in computed tomography

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electronic processing, thereby improving the utilization of the radiation dose. (2) The images can be displayed with enhanced contrast, thus overcoming the minimum contrast threshold of the human eye. (3) The CT reconstruction technique almost completely eliminates the superposition of anatomic structures, leading to a reduction of "structural" noise.

It is the random noise in a CT image that ultimately limits the ability of the radiologist to discriminate between two regions of different density. Because of its unpredictable nature, such noise cannot be completely eliminated from the image and will always lead to some uncertainty in the interpretation of the image. There is strong evidence that much of the random noise in present-day CT scanners is due to the statistical inaccuracies arising from the detection of a finite number of transmitted x-ray quanta. The properties of this statistical noise in CT images will be discussed in this chapter. The goal will be to characterize the noise content in a way that is closely related to the detection capabilities inherent in the image.

It will be shown in the discussion of the optimum receiver that the ability to detect a large-area object in a CT image is chiefly dependent on the NEQ, the total number of noise-equivalent quanta detected per unit length in the projections used to reconstruct that image. As such, it is not the fineness with which the x rays are detected in the projections (related to spatial resolution) but simply the effective total number of detected x rays that influences the large-area contrast sensitivity of the CT reconstruction. For a fixed, large-area contrast sensitivity, then, an increase in spatial resolution does not imply the necessity for higher dose! NEQ, which characterizes the low-frequency properties of noise in a CT image, can be determined from a CT noise image by measuring either the noise-power spectrum or the noise granularity function.

Since, for a given CT geometry, the number of transmitted x-ray quanta is proportional to the dose, the magnitude of the statistical noise depends on the dose. A figure-of-merit for dose utilization will be developed that can be used to compare the dose efficiencies of various CT scanners.

The x-ray computed tomographic (CT) scanner has made it possible to detect the presence of lesions of very low contrast (New et al., 1973). This dramatic improvement in detection capability over most conventional forms of x-ray imaging is a result of the following innovations: (1) the noise in the reconstructed CT images is significantly reduced through the use of efficient x-ray detectors and

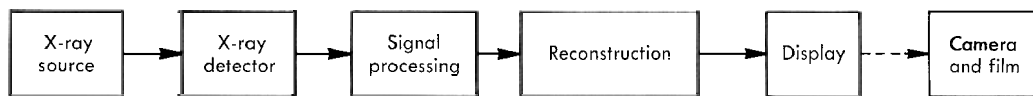


Fig. 113-1. Block diagram of the steps involved in obtaining a CT image.

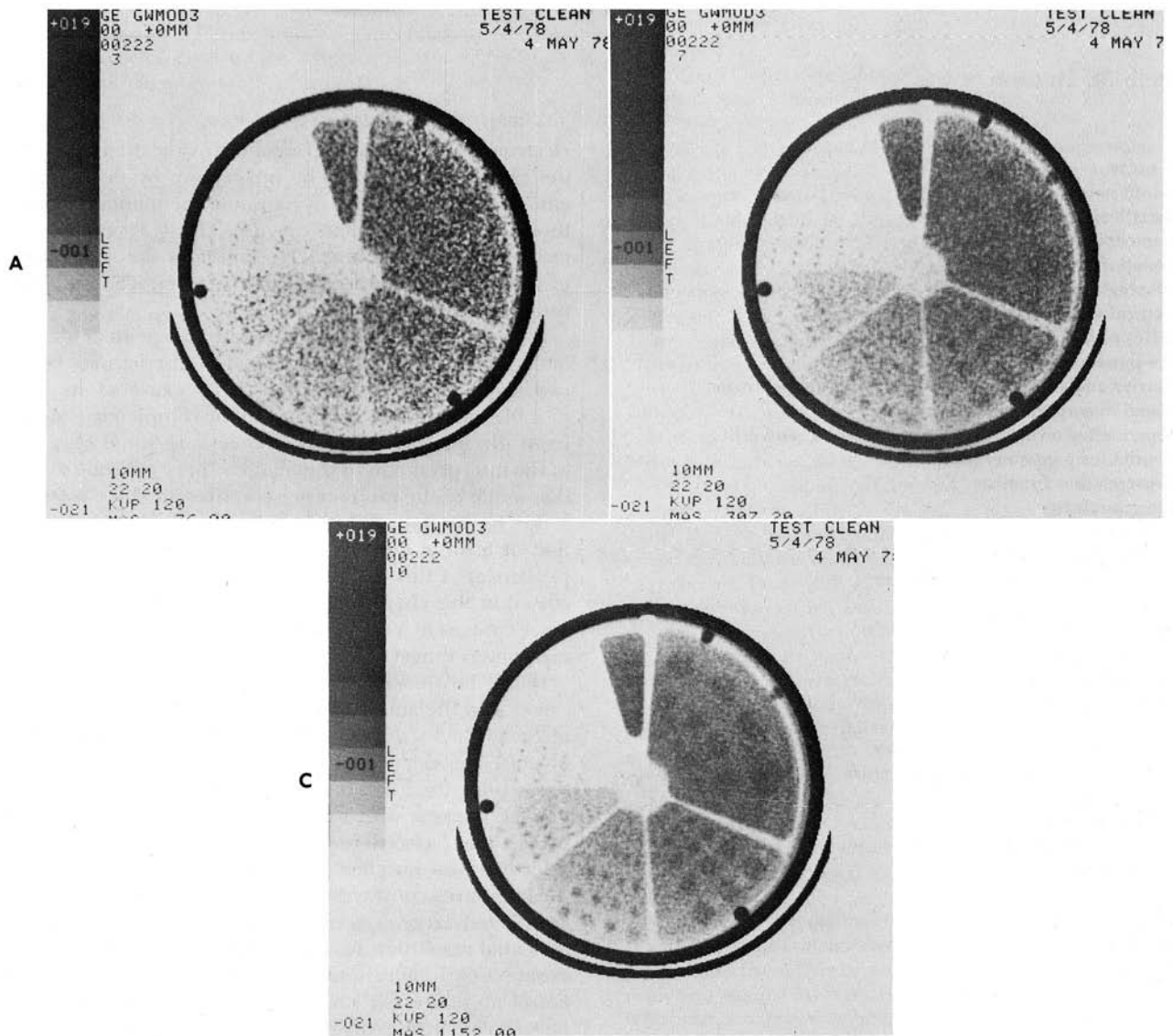


Fig. 113-2. Three scans of a contrast sensitivity phantom taken on a GE CT/T 7800 scanner at various milliampere-second values: A, 77. B, 307. C, 1152. The low contrast in the largest sector, 0.25%, is achieved by the partial volume effect. For this scanner higher dose leads to lower noise and improved contrast sensitivity. (Courtesy F. A. DiBianca, General Electric Co.)

IMAGE NOISE

There are several types of image "noise" that can interfere with the interpretation of an image. In the discussion of the types and sources of noise present in CT images it is helpful to consider the various steps in the production of a CT image as presented in Fig. 113-1. Although noise may infiltrate and corrupt the data at any point in the CT process, the ultimate source of noise is the random, statistical noise, arising from the detection of a finite number of x-ray quanta in the projection measurements.

Random noise

Image mottling, or fluctuations in the image density that change from one image to the next in an unpredictable and random manner, may be termed random noise. The radiologist is familiar with random noise in the form of radiographic mottle found in standard radiographs taken with fast screen-film combinations (Ter-Pogossian, 1967, p. 249).

Statistical noise

The energy in x-radiation is transmitted in the form of individual chunks of energy called quanta. Hence the response of an x-ray detector is actually the result of detecting a finite number of x-ray quanta. The number of detected quanta will vary from one measurement to the next, not because of inadequacies in the detection apparatus, but because of statistical fluctuations that naturally arise in the "counting" process. As more quanta are detected in each measurement, the relative accuracy of each measurement improves. *Statistical noise* in x-ray images arises from the fluctuations inherent in the detection of a finite number of x-ray quanta. Statistical noise may also be called quantum noise and is often referred to as quantum mottle in film radiography.

Statistical noise clearly represents a fundamental limitation in x-ray radiographic processes. The only way to reduce the effects of statistical noise is to increase the number of detected x-ray quanta. Normally this is achieved by increasing the number of transmitted x rays through an increase in dose. Fig. 113-2 shows a series of scans taken on a GE CT/T 7800 scanner* using a phantom designed to test contrast sensitivity. The reduction in the noise caused by increasing the x-ray exposure (milliamperere-seconds) is graphically demonstrated.

Electronic noise

In processing electric signals, electronic circuits inevitably add some noise to the signals. Analog circuits, those which process continuously varying signals, are most susceptible to additional noise. Thus the stage most likely to inject additional electronic noise in Fig. 113-1 is that of analog signal processing. The difficulty of noise suppression is compounded by the fact that for some types of x-ray detectors, the electronic signals are very small. There is evidence (Cohen, 1979) that many commercially available CT scanners are sufficiently well engineered to reduce the con-

tribution of electronic noise under normal operating conditions to a fraction of the statistical noise contribution. The signals are converted to digital or discrete form in the signal-processing step and then sent to a computer for reconstruction.

Digital circuits, those which process discrete signals as in digital computers, are relatively impervious to electronic noise problems.

Roundoff errors

Although digital computers are not subject to electronic noise, they do introduce noise in the reconstruction process through roundoff errors. The errors arise from the limited number of bits used to represent numbers in the computer. For example, the product of two numbers must be rounded off to the least significant bit used in the computer's representation of the number. Roundoff errors can normally be kept at an insignificant level either through choice of a computer with enough bits per word or through proper programming. It should be pointed out that in some CT scanners the final reconstruction is stored with the least significant bit equal to one CT number (0.1% of the linear attenuation coefficient of water). This should not influence the accuracy significantly so long as the rms noise is greater than one CT number.

Roundoff errors can and do occur in the display stage (Fig. 113-I), since CT display units can display only a fixed number of discrete brightness levels. As an example, a unit that has only 32 brightness levels will result in an increment size of about 0.05 optical density units (OD) on exposed radiography film. A difference of 0.05 OD produces a luminance contrast of 12%. Since under optimum conditions the human observer can discern objects at contrasts of less than 0.004 OD (Chapter 115; Burgess et al., 1979), such a unit can limit the observer's ability to interpret the film, particularly for wide display windows (100 or more CT units). This limitation may be circumvented by the use of a narrow window.

Artifactual noise

The wide variety of artifacts that can be produced by CT scanners is presented in Chapter 114. Artifacts might be viewed as a form of noise in that they interfere with the interpretation of the CT image. Their presence is often indicated by a readily identifiable pattern, for example, in the case of streak artifacts. These identifiable artifacts do not produce random noise, since they should be unchanged in repeated scans of the same object. However, there are instances in which regions of a reconstruction may experience an increase in variance due to nonapparent artifacts (Sheridan et al., 1980). Artifactual noise will not be considered further here.

Structural "noise"

Density variations in the object being imaged that interfere with the diagnosis are sometimes referred to as structural "noise" or structural clutter. In standard radiography a large amount of structural clutter is produced by the superposition of various anatomic structures, for example,

*General Electric Medical Systems Division, Milwaukee, Wis.

the image of rib bones overlaps that of the lung in a standard chest radiograph. The CT technique eliminates most of this superposition, but the radiologist should be aware that partial contributions may be introduced by structures that principally appear in adjacent CT slices.

Some organs, such as the liver, may have density variations within them that have the appearance of random noise. Although the texture pattern of the organ may not be reproducible from one CT scan to the next because of patient motion, this type of structural variation is, of course, not random. Indeed, the classification of this density variation as a type of noise is ill-advised, since the variation is intrinsic to the object itself. The study of the tissue texture may be interesting for its potential diagnostic value (Pullan et al., 1978).

PROPERTIES OF CT NOISE

The consequences of statistical noise in CT reconstructions have been discussed by numerous authors (Shepp and Logan, 1974; Cho et al., 1975; Huesman, 1975; Tanaka and Inuma, 1975, 1976; Barrett et al., 1976; Brooks and DiChiro, 1976; Chesler et al., 1977; Hanson, 1977; Huesman, 1977; Joseph, 1977, 1978; Hanson and Boyd, 1978; Riederer et al., 1978; Hanson, 1979a; Wagner et al., 1979). Several of these authors have pointed out that the process of reconstruction leads to some peculiar characteristics of the noise in CT images. The properties of statistical (quantum) noise in CT reconstructions will be explored in this discussion. Although the precise random noise pattern of any image cannot be predicted a priori, it is possible to characterize the average behavior of the noise by a variety of methods. Some of these methods give a complete description of the noise characteristics, such as the noise power spectrum or the noise autocorrelation function, whereas others give only a partial description, such as rms noise. It will develop that the noise fluctuation in one pixel of a CT reconstruction is *not* independent of the noise fluctuations in other pixels. Rather, the fluctuations in two separate pixels are, on the average, correlated.

Noise power spectrum

The noise power spectrum is a frequency representation of the correlations present in the noise. The ramplike nature of the noise power spectrum expected for CT reconstructions will be derived for the filtered back projection algorithm. It will be shown that noise power spectra of EM1 CT5005 reconstructions possess this ramplike behavior at low frequencies.

Fourier analysis - MTF and noise power spectrum

Any signal may be thought of as a sum of sine waves of appropriate frequency, amplitude, and position (phase). This decomposition of a signal into its frequency components is called Fourier analysis (see Chapter 110). It is often helpful to apply Fourier analysis to imaging systems (cameras, television, x-ray radiographs, etc.) because of the ease with which the imaging properties of various stages of the systems combine in their frequency representations. To be sure, Fourier analysis is only useful for linear systems, but

most imaging techniques are approximately linear for small changes in signal amplitudes. A good introduction to the application of Fourier analysis to radiography may be found in Johns and Cunningham (1969).

The modulation transfer function (MTF) is the frequency decomposition of a system's line spread response function.* The latter is the output of the imaging system when an extremely narrow line is used as input; for example, the radiograph produced of a very fine metal wire. Since in the frequency representation of an infinitely narrow line all frequencies are present with equal contribution, the MTF actually measures the frequency response of the system. Most MTFs are fairly close to unity at low frequencies, meaning that these frequencies are transferred through the imaging system with almost no degradation. The MTFs of all physical systems will eventually fall to zero at high frequencies. This means that the high frequencies are not preserved by the imaging system and simply are not present in the final image. Since perfectly sharp edges require infinitely high frequencies to represent them, the loss of the high frequencies results in blurred edges. Thus the MTF is the frequency representation of the sharpness of the system.

The noise power spectrum is related to the frequency decomposition of the image noise. The method of obtaining the noise power spectrum will be outlined to clarify the relationship. First, an image containing nothing but the noise to be analyzed is decomposed into its two-dimensional frequency components. These components, called the Fourier amplitudes, are then squared to obtain the "power" contribution at each frequency. This result is averaged with the results obtained from other similar images to arrive finally at the average noise power for each frequency, or, in other words, the noise power spectrum. The noise power spectrum is therefore proportional to the average (or mean) square value of the frequency amplitudes of the noise. Fig. 113-3 illustrates the computational procedure for obtaining the noise power spectrum of a one-dimensional noise sample. Note that the square of the amplitude is used instead of the amplitude itself, since the average value of the amplitude is zero while the average value of the squared amplitude is not. The noise power spectrum is often referred to as the Wiener spectrum. Often the response of the system is assumed to be circularly symmetric. Then the two-dimensional noise power spectrum may be reduced to a one-dimensional spectrum that is a function of only the radial frequency (distance from zero frequency in two-dimensions).

The noise found in conventional radiographs is typified by a noise power spectrum that is roughly constant over a wide range of frequencies. Such a noise power spectrum is called "white" in analogy to white light, which contains a mixture of light of all frequencies within the visible spec-

*The two-dimensional MTF is the frequency decomposition of a system's point spread function. The latter is the output of the imaging system when an extremely small point is used as input. A one-dimensional section of the two-dimensional MTF is obtained from the line spread function.

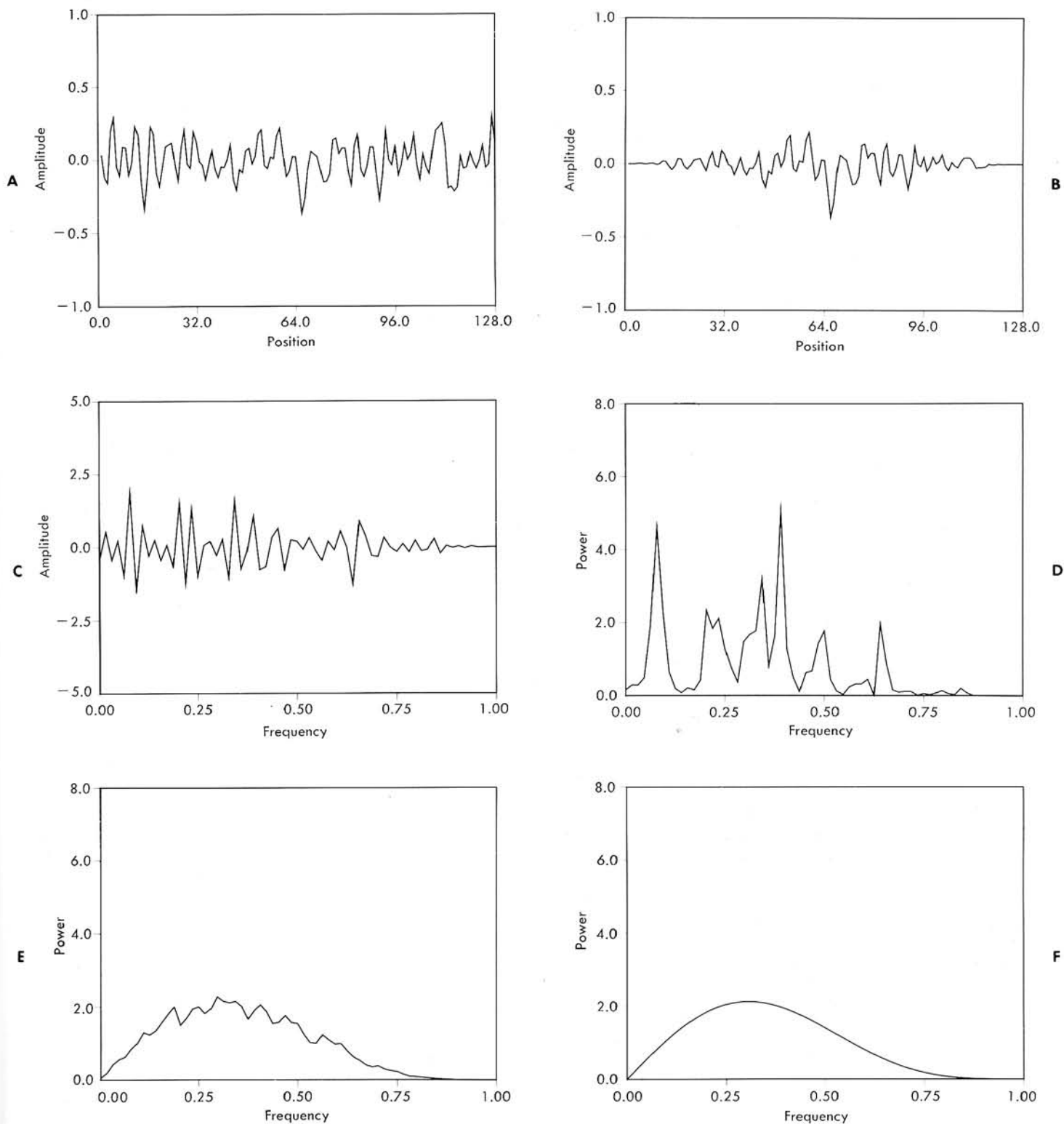


Fig. 113-3. Steps involved in the estimation of the noise power spectrum of a one-dimensional noise distribution. The first noise segment (A) is multiplied by a (Hanning) window to taper the edges to zero (B). The Fourier transform of B yields the Fourier amplitudes (C), which are squared to obtain the power spectrum of the first noise segment (D). The power spectra from 1000 noise segments are averaged to reduce statistical fluctuations in the estimated spectrum (E). If an infinite number of noise segments were employed, the resulting noise power spectrum would show no statistical fluctuations (F).

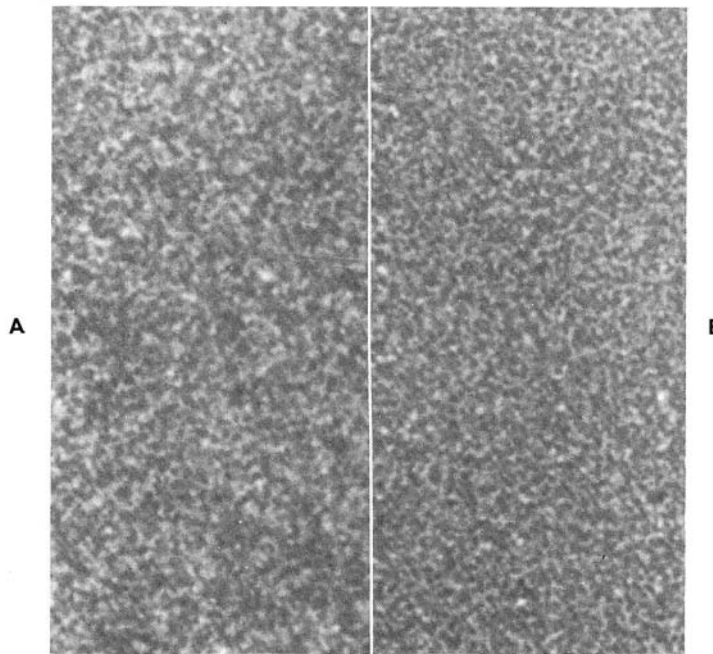


Fig. 113-4. Comparison of, A, "white" noise and, B, CT noise, each with the same rms deviation. The lack of low-frequency structure in the CT noise is evident, especially when viewed at a distance.

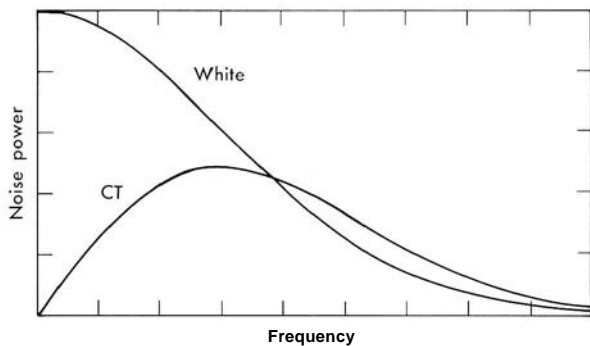


Fig. 113-5. Noise power spectra of the two images in Fig. 113-4.

trum. A white noise power spectrum indicates that the noise at one point in the image is uncorrelated with or independent from the noise at another point. In conventional radiographs this comes about since separate x rays are detected at two different points. Thus the statistical fluctuations giving rise to noise at one point are not related to the fluctuations occurring at a different position. Statements such as this can be made only if the points are separated by a distance sufficiently large to avoid "cross talk" between the points. Consequently, the noise power spectra of all physical systems must fall to zero at high frequencies just as the MTFs must do.

Visual example

The relationship between the noise power spectrum and the noise it characterizes may be displayed visually. Fig.

113-4 shows images containing two types of noise, one typifying white noise, and the other, CT noise. The rms deviation of the noise is the same in both images. The noise power spectra of these images is shown in Fig. 113-5. White noise, typical of film radiography (see above) is characterized by a flat power spectrum at low frequencies. In this example its power spectrum falls off at intermediate frequencies because of an assumed Gaussian MTF, that is, $\exp(-af^2)$. As discussed later, the CT noise power spectrum typically is linear at low frequencies going toward zero at zero frequency. In this example the CT noise power spectrum peaks at midfrequencies and then falls off with the same MTF as the white noise spectrum. The major difference between the two noise power spectra of Fig. 113-5 is the greatly diminished contribution at low frequencies typical of CT noise. Since low frequency corresponds to large distance, the CT noise image would be expected to have little contribution from clumps of noise with large area. The reduced blotchiness of the CT noise can be seen in Fig. 113-4, especially when the figure is observed from a distance so the eye is tuned to large-area structure.

Dependence on reconstruction algorithm and NEQ

The noise power spectrum for CT noise has been derived by Riederer et al. (1978). The derivation is based on the filtered back projection algorithm (Shepp and Logan, 1974) for a parallel-beam geometry in which a constant density of x rays is detected in all the projections. The corrective filter that is applied to the Fourier transforms of the projections is assumed to be:

$$G(f) = |f| H(f) \quad (113-1)$$

where f is the frequency and H is a weighting (or apodization) factor. The principal restriction on H is that it be nearly unity for very low frequencies f . It may be shown that the noise power spectrum S for statistical CT noise can be written as follows (Riederer et al., 1978; Hanson, 1979a; Wagner et al., 1979):

$$S(f) = \frac{\pi |f| H^2(f)}{\text{NEQ}} \quad (113-2)$$

where NEQ, the number of noise-equivalent quanta, is the total effective number of x-ray quanta detected per unit distance along the projections (summed over all the projection measurements). If there is no source of noise other than statistical and the reconstruction algorithm is efficient, then NEQ will be just the total number-of detected x rays per unit projection length. The presence of other sources of noise can reduce NEQ compared with the actual number of detected quanta. Also, the measurement of the x-ray fluence by means of an energy-integrating detector (as is done on essentially all commercial CT scanners) leads to a slight reduction (-10%) in NEQ compared to that attainable through x-ray counting. In equation 113-2 it is assumed that the reconstructed image is the linear attenuation coefficient μ given in units of (cm^{-1}). Then S will be dimensionless and NEQ will have dimension (cm^{-1}), as it should.

Equation 113-2 indicates that the noise power spectra of CT reconstructions should have a ramplike dependence at low frequencies (where $H = 1$). The slope of S at low frequencies is determined by NEQ. As such, NEQ may be used to characterize the large-area, low-contrast detection capabilities of CT images (see discussion of optimum receiver). It shall be shown that equation 113-2 implies there is no loss in detection performance in a filtered back projection reconstruction compared to the input projection data for large objects. As more quanta are detected, S will be reduced inversely and the CT image noise will diminish.

The dependence of S on the weighting factor H of the algorithm should be noted. $|H|$ may be regarded as the MTF associated with the reconstruction algorithm in the absence of binning problems associated with the finite size of the reconstruction pixels (Hanson, 1979b). The MTF of the complete scanner system is the product of all the MTFs that contribute to the spatial resolution of the final image. For many CT scanner systems the dominant contribution is the MTF associated with the finite width of the radiation beam used in the projection measurements (Barnes et al., 1979). Thus, although the total noise of the CT scanner is closely related to the spatial resolution of the reconstruction algorithm, it may have only a weak dependence on the overall spatial resolution. It should also be realized that the image noise is not dependent on the pixel size except insofar as the latter affects the choice of H .

Results for a commercial scanner

The noise power spectrum has been calculated for an EMI CT5005 scanner* to demonstrate that it has the form

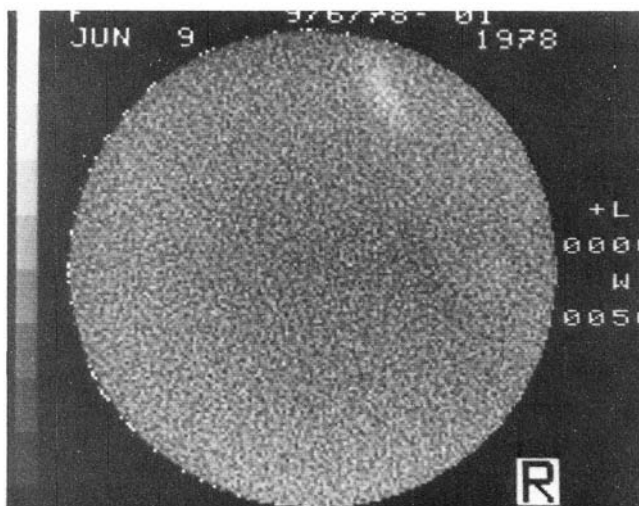


Fig. 113-6. EMICT5005 scan of the EMI water phantom used in calculation of noise power spectrum. The light spot near the top of the reconstruction is a roller mark produced by the film-processing unit. (Courtesy D. P. Boyd, University of California, San Francisco.)

predicted by equation 113-2. The EMI 26 cm diameter water calibration phantom was scanned six times at 140 kVp, 28 mA using the normal scan time (20 seconds). The 320×320 reconstruction matrix was obtained by decoding the EMI magnetic tape. The pixel size is 0.75 mm. These scans were performed at the University of California (San Francisco) Medical Center in October, 1977, and the scanner was not equipped with the new dose-reducing collimator (EMI 5221).

Analysis of the reconstructions showed that the rms deviation (see equation 113-3) of the noise (1.28%) was constant over the whole image to very good accuracy. Furthermore, the average values in the reconstructions were uniform to better than 0.1%. Fig. 113-6 shows one of the water scans used. Any variations seen in Fig. 113-6 can arise only from either the display unit or subsequent film processing. The presence of the film-roller mark demonstrates the need for quality control in these often neglected aspects of CT technology.

The noise power spectrum of the EMI reconstructions was calculated in a manner similar to that described in the earlier discussion of noise power spectrum. The resulting spectrum, Fig. 113-7, has the predicted linear behavior at low frequency (below 0.1 mm^{-1}). The slope of the spectrum at midfrequencies (0.1 to 0.3 mm^{-1}) is greater than at low frequency, indicating that EMI has used a value of H which is greater than 1 at these frequencies. The resulting edge enhancement sharpens edges in the reconstruction to a slight degree. Such edge enhancement is usually accompanied by overshoots and undershoots in the step response function.

The NEQ of the EMI scans can be obtained from Fig. 113-7 by using equation 113-2. The result is $\text{NEQ} = (1.85 \pm 0.03) \times 10^7 \text{ mm}^{-1}$. Thus the total number of equivalent x-ray quanta detected in each scan (product of

*EMI Medical, Inc., 4000 Commercial Dr., Northbrook, Ill. 60062.

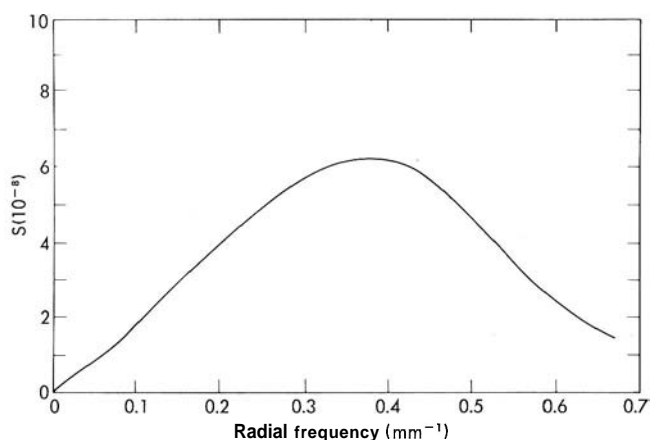


Fig. 113-7. Noise power spectrum for normal 20-second scans on an EMI CT5005 of a 24 cm diameter water phantom. The slope of the spectrum at low frequencies determines NEQ.

NEQ times the length of each projection) was $260 \text{ mm} \times 1.85 \times 10^7 \text{ mm}^{-1} = 5 \times 10^9!$

As was pointed out previously, there are numerous sources of noise other than statistical that may contribute to CT reconstructions. These other noise sources may not have noise power spectra similar to that expected for statistical noise. It has been verified that the noise power spectrum of the EMI Mark I has the expected form (Wagner et al., 1979). However, it remains to be determined whether other CT scanners produce images with similar noise power spectra.

Autocorrelation function

The autocorrelation function provides a measure of the correlation between noise fluctuations that occur at two different points in the image. In general, it is a function of the spatial distance between the points, both in X direction Δx , and Y direction Δy . It is thus two-dimensional in nature. If the autocorrelation depends only on the distance between point pairs, $d = \sqrt{(\Delta x)^2 + (\Delta y)^2}$, then it is called isotropic. The autocorrelation for a separation Δx in the X direction and no separation in the Y direction is calculated as follows: (1) the reconstructed value at a point x, y , $r(x, y)$, is multiplied by the value at $x + \Delta x, y$, $r(x + \Delta x, y)$, to form the product $r(x, y) r(x + \Delta x, y)$; (2) the autocorrelation function is determined by averaging this product over all values of x and y throughout the image. A similar procedure is followed when a y separation between the point pairs, Δy , is desired.

When the autocorrelation of the noise is positive for a given separation, d , the noise is said to be positively correlated at that distance. This means a fluctuation at a particular point in the image is, on the average, going to be accompanied by a fluctuation of similar sign at a point a distance d away. Similarly, a negative autocorrelation function for a separation d means that a fluctuation at one point is, on the average, accompanied by a fluctuation of opposite sign at a point a distance d away. The noise is said to be

uncorrelated at a distance d if the autocorrelation function is zero for that distance.

The autocorrelation function of the noise is intimately related to the noise power spectrum. It is just the inverse two-dimensional Fourier transform (FT) of the noise power spectrum. In fact, the most economical way of estimating the autocorrelation for a given noise image is to first estimate the noise power spectrum and then calculate its inverse FT (Wagner, 1976). Since the autocorrelation function may be calculated from the noise power spectrum, and vice versa, both of these methods provide equivalent and complete descriptions of the correlations present in the noise.

The autocorrelation function can provide some enlightenment as to the fundamental differences between white and CT noise. The autocorrelation function corresponding to white noise is zero for all separations greater than some minimum distance. (This can be related to the flat noise power spectrum that defines white noise, since the inverse FT of a flat spectrum is simply a δ -function, that is, non-zero at only one position—zero separation.) This indicates that for white noise, the noise fluctuation at one point is not correlated with the fluctuations that may occur at any other point. The value at each point in the image fluctuates in a truly independent and random fashion.

For CT noise, however, the ramp behavior of the noise power spectrum at low frequencies implies a long-range negative correlation of the noise (Hanson, 1979a). Thus the value at each point in the image is correlated or influenced to some degree by the noise values at all other points. If a strong positive noise fluctuation exists at one point, there will be a tendency for the noise fluctuations at nearby points (but beyond the spatial resolution of the reconstruction algorithm) to be negative. This negative correlation decreases as the separation of the points increases.

Noise granularity

Several measures of the granular nature of image noise have been used in the past to investigate film noise (Dainty and Shaw, 1974). Although it is not as fundamental as the noise power spectrum, the noise granularity emphasizes some important properties of CT noise. The noise granularity displays the graininess of the noise as a function of the area over which the noise is averaged. It is defined in such a way that white or uncorrelated noise leads to a constant granularity independent of the area. It will be shown that the long-range negative correlations present in CT noise lead to decreasing granularity as the area increases. The noise granularity can be calculated from the noise power spectrum, or it can be estimated directly from the noise image.

Definition

The rms deviation from the mean, usually represented by σ , of a set of N values $V_1, V_2, V_3, \dots, V_N$ is defined as:

$$a = \text{rms deviation} = \frac{\sqrt{(V_1 - \bar{V})^2 + (V_2 - \bar{V})^2 + \dots + (V_N - \bar{V})^2}}{N - 1} \quad (113-3)$$

where the average or mean value \bar{V} is:

$$\bar{V} = \frac{V_1 + V_2 + V_3 \dots + V_N}{N} \quad (113-4)$$

The rms deviation σ is a measure of the amount of scatter or variation in the values away from their mean \bar{V} . Suppose the mean value of the reconstruction is calculated for an area A at N locations throughout the image. Then the noise granularity shall be defined here as:

$$G(A) = \sigma_A \sqrt{A} \quad (113-5)$$

where σ_A is the rms deviation (equation 113-3) of these means and A is the area. Thus the noise granularity is a measure of the amount of scatter in the area-averaged values. G , defined by equation 113-5, is Selwyn's granularity coefficient (Selwyn, 1935) divided by $\sqrt{2}$.

It is convenient to extend the definition of noise granularity to include weighted means. The weighted mean is calculated by multiplying the reconstruction values by a weight function and then summing (or integrating) the resulting products. In equation 113-5, σ_A is then the rms deviation of the weighted means. The weight function is a function of x and y , which usually is nonzero only within a certain distance from the central pixel. It is conveniently normalized such that the sum of the weights is unity, and it is characterized by an effective area A . For the strict unweighted mean the weight function is a constant over the area A and zero elsewhere. The weighted mean then yields the average value over the area A . A more interesting weight function is, for example, pyramidal in shape. The largest weight is given to the central pixel, and the weight function drops linearly to zero in all directions away from the center. The application of this weight function will be demonstrated later.

Relationship to noise power spectrum

G may be readily calculated in terms of the noise power spectrum S , since σ_A^2 is simply the integral over frequency of S multiplied by an appropriate function (the square of the frequency representation of the weighting function used in calculating the mean). Thus S should be considered as a more fundamental measure of the noise characteristics than G . However, there may be some practical advantages, such as ease of calculation, to the use of G . Moreover, G is useful because it demonstrates the effect on the rms noise (σ_A) of averaging over different areas.

The relationship between G and S may be used to show that G for large areas is principally determined by the magnitude of S at low frequencies. For CT noise the ramplike dependence of S at low frequency, or equivalently the long-range negative correlation in the noise, implies that (Hanson, 1979a):

$$G(A) \approx \frac{1.25}{A^{1/4} \sqrt{NEQ}} \quad (113-6)$$

where NEQ is the number of equivalent x-ray quanta detected per unit length in the projections (equation 113-2). It should be noted that equation 113-6 is a good approximation for weighting functions that have tapered edges,

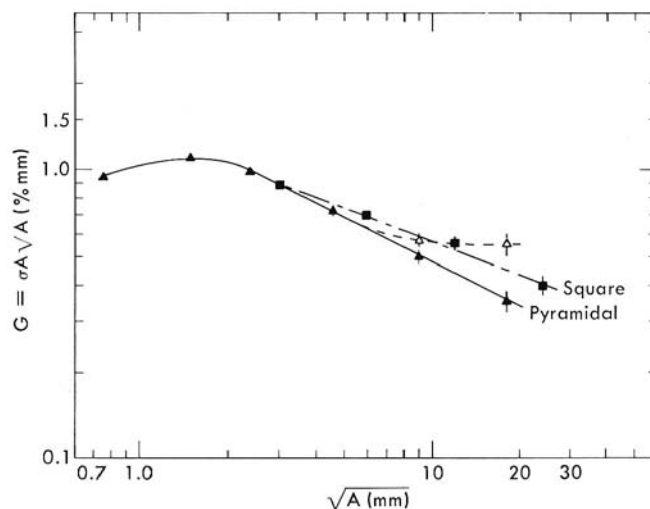


Fig. 113-8. Noise granularity for the EM1 CT5005 scans. The square data points were calculated using constant weight throughout square regions, that is, unweighted means. The triangular data points result when a pyramidal weighting function is used. The open triangles show the effect of not removing frequencies below 0.02 mm^{-1} .

such as Gaussian and pyramidal functions. Equation 113-6 does not hold when G is calculated for unweighted means. The NEQ for a particular CT scanner may be determined from a careful measurement of $G(A)$ using tapered weighting functions (Hanson, 1979a).

For CT noise, equation 113-6 indicates that $G(A)$ will decrease slowly as the area A increases ($G(A) \propto A^{-1/4}$). This is interesting, since for white noise, $G(A)$ is independent of the averaging area, A . One interpretation of this effect is that as the averaging area increases, the reduction in rms deviation of CT noise is hastened by the negative correlations in the noise.

Results for a commercial scanner

The noise granularity has been calculated for the same EM1 CT5005 water bath scans described previously. As shown in Fig. 113-8, G falls off for large A . With the pyramidal weighting function, the falloff is in accordance with equation 113-6. The NEQ obtained from this curve using equation 113-6 is the same as that obtained from the noise power spectrum. The curve derived from unweighted (square) means does not drop as fast. The sharp edges of the square weight function generate substantial high-frequency components that accentuate the high-frequency portions of the CT noise power spectrum.

The solid data points in Fig. 113-8 were obtained after the noise images were flattened by removal of the very low frequencies. When this is not done, the granularity does not continue to drop at large areas (open triangles) even though the uniformity of the initial EM1 reconstructions was good to begin with. Earlier calculations of the granularity (Boyd et al., 1976, 1977), which did not show the falloff of the granularity at large areas, apparently suffered from a lack of uniformity in the reconstructions.

rms noise

Frequently the rms deviation (equation 113-3) of the noise is quoted for CT reconstructions. It should be clear that the rms noise is not a complete characterization of CT noise, since it ignores the frequency dependence of the noise. Specifically, the rms noise depends critically on the weighting factor $H(f)$ used in the reconstruction algorithm, which will vary from one CT scanner to the next. It will be shown that the ability to detect large-area objects is related to the low-frequency noise power content, which is relatively unaffected by $H(f)$. Thus detectability is not simply related to the rms noise.

The rms noise may be calculated from the noise power spectrum (by Parseval's theorem) as the square root of the total noise power:

$$\sigma = \sqrt{\iint df_x df_y S(f_x, f_y)} = \sqrt{2\pi \int df f S(f)} \quad (113-7)$$

Since σ^2 is the integral of the noise power spectrum, it contains no information about the frequency dependence of the noise.

CT NOISE AND DOSE

The statistical noise in CT reconstructions is closely related to the dose. The dose efficiency of a given scanner may be obtained by comparison of its delivered dose to that required by an ideal x-ray scanner to achieve the same NEQ. This dose efficiency summarizes all contributions to the utilization of the delivered dose such as x-ray beam collimation, compensation wedges, detector efficiency, and signal processing, as well as the degrading effects of polychromaticity, detection of scattered radiation, electronic noise, and inefficient reconstruction algorithms.

Dose dependence

The dependence of the statistical noise power spectrum on dose is indicated in equation 113-2. Since the number of detected quanta (NEQ) is proportional to the number of initial quanta, and hence the dose, D , it is concluded that S is inversely proportional to dose. Then, by equation 113-7, the rms value of the statistical noise should be inversely proportional to \sqrt{D} . This is not always the case for commercial CT scanners.

Fig. 113-9 shows the dose (proportional to mAs) dependence of the rms noise in reconstructions from a GE CT/T 8800 scanner of an 18-cm-diameter phantom (Cohen and DiBianca, 1979). The $(D)^{-1/2}$ dependence expected for statistical noise is represented by the dashed line. The departure of the measured noise from this dependence indicates the presence of contributions to the total noise other than statistical. The source and frequency characteristics of these background noise contributions have not been determined. For the GE CT/T 8800, the rms noise is proportional to $D^{-1/2}$ up to about 200 mAs, indicating that at the lower exposures the noise may be statistically limited.

A word of caution must be given. Other sources of noise may also follow the $D^{-1/2}$ dependence, depending on how the increase in dose is achieved. If, for instance, the dose is increased by simply integrating the x-ray signals over

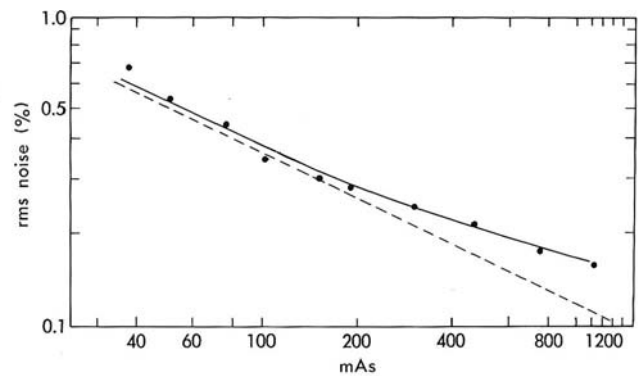


Fig. 113-9. The rms noise dependence on exposure (milliampereseconds) for the GE CT/T 8800 scanner. The dashed line shows the dependence expected for statistical noise. (Modified from Cohen, G., and DiBianca, F. A.: *J. Comput. Assist. Tomogr.* **3**:189-195, 1979.)

longer time intervals or by increasing the scan time, other random noise sources (e.g., electronic) may be reduced in the same way as statistical noise, leading to the same $D^{-1/2}$ behavior. Thus, although the $D^{-1/2}$ dependency is *necessary*, it is *not sufficient* to prove that the scanner noise is "quantum limited." Proof must be based on the comparison of the measured NEQ to the actual number of x-ray quanta detected.

Dose efficiency

The dose required to obtain a given NEQ for any practical scanner will always be more than that required for an ideal scanner. How much more will depend on scanner design, principally in regard to x-ray beam filtration and collimation, and detector quantum efficiency. A good measure of dose efficiency is:

$$\eta = \frac{D_{NEQ}}{D} \quad (113-8)$$

where D_{NEQ} is the minimum dose required to produce the measured NEQ for the phantom used and D is the actual dose delivered by the scanner.

Dose specification

The dose D may be specified in several ways. Common current practice is to specify the peak surface dose for either a single scan or a series of scans. Of course, the surface dose is easiest to measure, since the phantom surface is readily accessible. However, the specification of the dose as the average dose throughout the whole slice for a series of scans has both calculational and physiologic advantages. The average dose is an appropriate measure of physiologic damage or risk if radiation damage effects are proportional to integral dose. The measurement of the average dose is more difficult, and several dose measurements are required over the cross section of the phantom.

D_{NEQ}

D_{NEQ} will be calculated for a highly idealized scanner. It is assumed that a perfectly collimated, monochromatic

x-ray beam is employed. The dose is calculated for a cylindrical phantom with no additional bolus material surrounding it. The x-ray source intensity is varied to maintain a constant unscattered intensity density at the detector (Fig. 113-10). Hence NEQ is held fixed across the projection and the rms noise is constant throughout the resulting reconstruction. The unscattered beam is detected with 100% quantum efficiency with no contribution from scattered radiation. The average dose in a series of scans is:

$$D_{\text{NEQ}} = \frac{\text{Deposited energy}}{\text{Mass}} = \frac{E N_{\text{TOT}} \tilde{\mathcal{F}}}{\rho V} \quad (113-9)$$

where N_{TOT} is total number of x rays incident on the phantom, E is x-ray source energy (monochromatic), V is volume of each slice, ρ is density of the phantom, and $\tilde{\mathcal{F}}$ is fraction of the total incident energy that remains in the phantom. Under the conditions just described the average dose is:

$$D_{\text{NEQ}} = \frac{4E \text{ NEQ} \tilde{\mathcal{F}}}{\pi \rho d t} I(\mu d), \quad (113-10)$$

where d is phantom diameter, t is distance between adjacent scans, and I^* is linear attenuation coefficient at energy E . The integral I required to relate N_{TOT} to NEQ is:

$$I(\mu d) = \int_0^1 e^{\mu d \sqrt{1-x^2}} dx \quad (113-11)$$

$$\approx \frac{e^{\mu d}}{\sqrt{\mu d}} \left[1.255 - \frac{0.515}{\mu d} \right] \quad (113-12)$$

The approximation given in equation 113-12 is accurate to within 0.5% for $\mu d \geq 2$, which applies to most medical situations.

It will be noted that D_{NEQ} (equation 113-10) is inversely proportional to the scan separation, t . To maintain a fixed noise level (NEQ), the dose has to be increased when the slices are taken at smaller separations. The choice of t may be viewed as similar to the choice of the dose, which should be made on the basis of the diagnostic task at hand.

The fraction of energy that is backscattered from a planar phantom (Berger and Raso, 1960) places an upper limit on $\tilde{\mathcal{F}}$ of about 0.75 in the CT diagnostic range of 60 to 100 keV. It is estimated that a further 10% to 25% of the incident energy will scatter out of or be transmitted through waterlike phantoms with diameters between 20 and 40 cm. $\tilde{\mathcal{F}}$ is estimated to be approximately 0.6 for a 20 cm diameter phantom and should be only weakly dependent on the diameter.

The strong dependence of D_{NEQ} on the phantom diameter is evident in equations 113-10 and 113-12. Consequently, it is important that the phantom diameter be stated along with measurements of NEQ or dose. On the other hand, the dose efficiency η should have little dependence on d .

It should be borne in mind that the dose efficiency calculated here is based on a comparison to an ideal CT scanner that employs a monochromatic x-ray beam. All practical scanners are restricted to the use of polychromatic x-ray sources. The utilization of polychromatic x rays leads

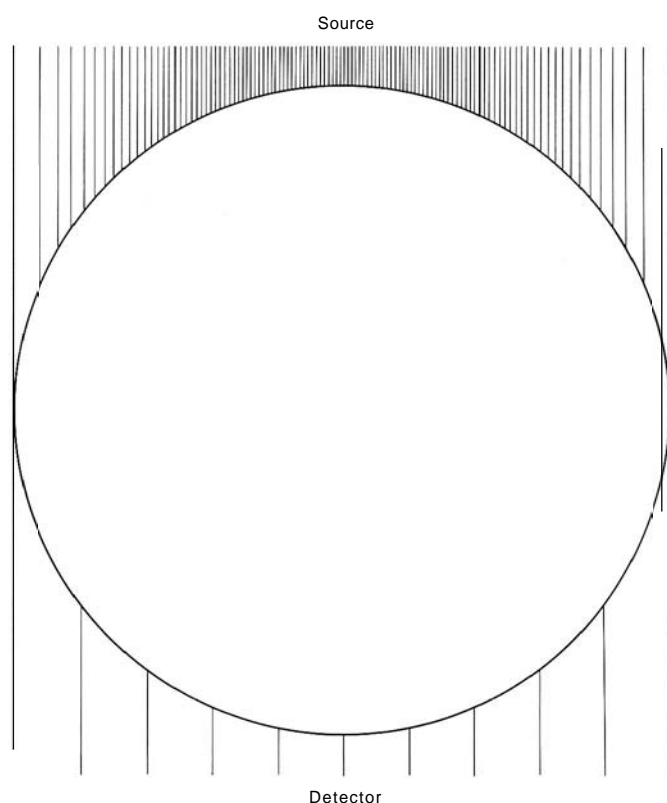


Fig. 113-10. Idealized x-ray geometry assumed for calculation of D_{NEQ} . The source intensity is reduced at the edges of the phantom to maintain a constant number of unscattered quanta per unit length of detector.

to a reduction in η for two reasons. The first is that with the light beam filtration often used, a significant portion of the deposited dose arises from low-energy x rays that contribute little to the detected signals (because of enhanced absorption in the phantom). The second reason is that for the detection schemes employed, the detector response is proportional to the total x-ray energy rather than to the number of x-ray quanta. This typically reduces η by about 10%.

Results for a commercial scanner

As an example, the dose efficiency of the EMI CT5005 scanner may be estimated. The average dose for a normal scan of the 10-inch calibration phantom is estimated to be about 3.7 rad for a series of scans taken at 1 cm intervals. E is set equal to the effective energy of the scanner, 82 keV (Millner et al., 1978). If $d = 26$ cm, $\mu = 0.182 \text{ cm}^{-1}$ (Storm and Isreal, 1970), and $\tilde{\mathcal{F}} = 0.6$, the dose required to achieve an NEQ of $1.85 \times 10^7 \text{ mm}^{-1}$ with an ideal scanner is $D_{\text{NEQ}} = 0.43$ rad. Thus the dose efficiency of the EMI CT5005 is about $\eta = 12\%$ compared with the ideal scanner. Note that this efficiency is quoted for a CT5005 scanner not equipped with the now-available dose-reducing collimator. It should be noted that the η calculated here includes all contributions to the dose efficiency. It is a far more stringent mea-

sure of dose efficiency than just the "detector efficiency."

DETECTABILITY IN THE PRESENCE OF CT NOISE

It is clear from Fig. 113-2 that a reduction in the magnitude of the noise leads to improved detection of contrast differences. But suppose the same phantom were scanned on another scanner with completely different spatial resolution and, more important, a completely different reconstruction algorithm. What measure of noise would allow comparison of the ability to discriminate contrast differences in the two scans? This is the central issue in the description of noise. It is important to use parameters to characterize the noise that are directly and simply related to detectability. In this discussion the close relationship between NEQ and the detection of large-area objects is described. This relationship is established for the optimum receiver, which fully takes into account the characteristics of CT noise. The connection between the detection performance of the optimum receiver and that of the human observer is not well documented. However, it is expected that the two performances will track each other in a relative evaluation of similar images. Thus the basis of comparison dictated by the optimum receiver will probably be useful in the comparison of clinical images.

Detection task

To simplify matters, only the binary decision problem will be considered here. The decision to be made is whether a specific object is present at a specific location. Furthermore, it is assumed that the background on which the object is superimposed is completely specified. This detection problem is exemplified by the phantom in Fig. 113-2 in which circular objects are present on a flat background. However, the presence of a row of circles rather than a single circle alters the problem slightly and complicates the analysis of the results. The binary decision case may be extended to the multiple decision problem (Goodenough, 1975; Goodenough and Metz, 1974) or to the problem of the search for objects within an image (Wagner, 1977). Although the binary decision problem represents a gross simplification of the clinical detection situation, at present this simplification is necessary to permit theoretical analysis and psychophysical testing.

Clinical diagnosis clearly relies heavily on the ability of the radiologist to recognize patterns. The general pattern recognition problem is very difficult to model in full detail. However, the ability to detect component parts of a pattern must form the basis of pattern recognition. Thus it is hoped and expected that results obtained from analysis of the simple detection problems often encountered in psychophysical testing will be relevant to the more complex clinical situation.

Optimum receiver

Detection sensitivity index, d'

Given an object to be detected in the presence of a specific type of noise, the best detection performance that is possible may be determined through application of signal

detection theory (Whalen, 1971; Van Trees, 1968; Wagner, 1978). The best decision criterion that can be used in a given detection problem is referred to as the "optimum receiver" in signal detection theory. The optimum receiver will depend on the situation at hand. In particular, the optimum receiver must take into account the properties of the noise to be "optimum." It is often possible to characterize the detection performance of the optimum receiver without actually constructing or implementing the detection criterion.

The detection performance of any detector applied to a given detection task may be summarized by its receiver operating characteristic (ROC) curve. The ROC curve is a plot of the probability of a "true positive" response versus the probability of a "false positive" response (Chapter 115; Green and Swets, 1966). For additive, Gaussian distributed noise, the ROC curve for the binary decision problem may be completely specified by a single parameter, the detection sensitivity index d' . The d' depends on the object's contrast, size, and shape as well as on the magnitude and correlations of the noise. For the optimum receiver d' may be expressed in terms of the frequency representation of the object $R(f)$ as follows (Barnard, 1972):

$$d'_{\text{OPTIMUM}} = \sqrt{\iint df_x df_y \frac{|R(f)|^2}{S(f)}} \quad (113-13)$$

where S is the noise power spectrum (see earlier discussion of dependence on reconstruction algorithm and NEQ). It is observed that d'_{OPTIMUM} is determined by the frequency sum or integral of the ratio of the signal power to the noise power. It should be noted that the design of an optimum receiver depends critically on the properties of the noise. Thus a receiver that is optimum for white (uncorrelated) noise will not be optimum for CT noise.

The plausibility of equation 113-13 may be illustrated for two limiting cases. For the first case the following situation will be considered: S is zero for some finite frequency interval in which the object power $|R|^2$ is not zero. The integrand in equation 113-13 would then be infinite over that frequency interval yielding an infinite value for d'_{OPTIMUM} . This is reasonable, since the optimum receiver would only have to check the image power (after the known background was removed) in the appropriate frequency interval. If there was any power present, it could only be due to the object. The optimum receiver could never make a mistake! Hence, $d'_{\text{OPTIMUM}} = \infty$. In the second case, the situation in which $|R|^2$ is zero over some finite frequency interval is considered. Equation 113-13 indicates that noise power in that frequency interval will not influence the optimum receiver. Again, this is reasonable, since the optimum receiver can remove these frequency intervals from consideration by Fourier transformation of the image followed by zeroing out the Fourier amplitudes in the relevant frequency interval.

Equation 113-13 leads one to an interesting conclusion concerning the trade-off between noise magnitude and spatial resolution. It is well known that the rms noise may be reduced by smoothing the image. Smoothing also results in a loss of spatial resolution, which is supposed to make it more difficult to detect small objects or to locate the posi-

tions of sharp edges. However, image smoothing is equivalent to the multiplication of the frequency representation of the image by a filter, which generally reduces the high-frequency components of the image. Since both $|R|^2$ and S are affected by the filter in the same way, equation 113-13 indicates that d'_{OPTIMUM} is not altered by the smoothing process, unless the filter is zero for some finite range of frequencies where $|R|^2$ is not zero. Thus the performance of the optimum receiver is not affected by smoothing (unless information is lost by a zero filter). Indeed, it is not necessary to trade off between low noise and high spatial resolution for the optimum receiver.

These statements concerning the optimum receiver may or may not have bearing on what might be expected of a human observer. For example, the human observer may suffer critical band masking between frequency intervals (see Chapter 115). Thus it is unlikely that the human observer could make use of information in one frequency interval that through filtering was reduced by a factor 100 relative to neighboring frequency intervals. Of course, the "optimum receiver," being a conceptual entity, would have no difficulty recouping the information in the attenuated frequency band.

Although the optimum receiver may not realistically characterize the performance of the human observer, it provides the ultimate standard against which the human observer may be compared. If it is found that the performance of the human observer falls short of this ideal in the simple detection task envisioned, then it may prove useful to explore the reasons for the shortcomings of the human observer.

Application to CT

The application of signal detection theory to computed tomography leads to an interesting result (Hanson, 1979a). Of course, the projection data themselves may be used to detect the presence of an object within the projection field. It is found that when the totality of projection data is analyzed by the optimum receiver, the resulting d'_{OPTIMUM} is equal to or greater than the d'_{OPTIMUM} obtained from analysis of the reconstruction. In other words, the detection in a CT reconstruction of a specific object on a known background can be no better than when the detection of the same object is based on the direct projection measurements. Furthermore, the detection performance based on an efficient reconstruction can equal that based on the projections. It has been shown that the filtered back projection algorithm is efficient, in this sense, for the detection of large objects (Hanson, 1979a, 1980). In the practical case of reconstruction in a discrete pixel array from discretely sampled projections, there can be a loss of information leading to some degradation in detection sensitivity (Hanson, 1979b).

The frequency representation, $R(f)$, of an object with large area is concentrated at low frequencies. Then equation 113-13 indicates that the detection sensitivity of large-area objects will be principally determined by the low-frequency content of the noise power. Since statistical CT noise has a ramplike noise power spectrum at low frequencies, the single parameter that characterizes the slope

of the ramp, NEQ, is a sufficient measure of the detection sensitivity for large objects. It is found (Hanson, 1979a) that the optimum sensitivity index is:

$$d'_{\text{OPTIMUM}} = \Delta\mu A^{3/4} \sqrt{\text{NEQ}} \quad (113-14)$$

where $\Delta\mu$ is the average contrast of the object with an effective area A . Equation 113-14 is a good approximation for most objects of large area (square, circle, etc.). The $A^{3/4}$ dependence of d'_{OPTIMUM} should be noted. It arises from the ramplike nature of the CT noise power spectrum. For white noise ($s = \text{constant}$), d' is proportional to $A^{1/2}$.

Example

An example can illustrate the use of equation 113-14. Consider a 5 mm diameter cylinder placed in a water phantom with a 26 cm diameter. What is the value of d' achieved by an EM1 CT5005 scanner if the reconstructed density of the cylinder differs from water by 0.4%? Since $\mu = 0.019 \text{ mm}^{-1}$ for water, and the area of the reconstructed circle is $A = \pi \times 2.5^2 = 19.6 \text{ mm}^2$, equation 113-14 gives:

$$d' = 0.004 \times 0.019 \times (19.6)^{3/4} \sqrt{1.85 \times 10^7} = 3.0 \quad (113-15)$$

For the binary decision problem, a d' of 3.0 results in an ROC curve such that a true positive response can be made 93% of the time for a false positive probability of 7%. If the position of the cylinder is unknown, then its detection is made much more difficult by the necessity of a search procedure. The resulting ROC curve would be much poorer than that for the binary decision case (Goodenough and Metz, 1974).

Human observer

The relationship between the detection capabilities of the ideal detector and those of the human observer has not been fully explored for images containing CT noise. It is possible that human observers may have shortcomings, particularly in their ability to integrate the noise over the object area. The unusual correlations present in CT reconstructions may prove difficult for the eye-brain to take into account. Several psychophysical studies (Hanson, 1977; Joseph, 1977, 1978; Chew et al., 1978; Orphanoudakis*) have shown that under certain circumstances observer detectability of large objects is improved by smoothing CT images. The reason for this improvement remains to be explained. Furthermore, the $A^{-3/4}$ dependence in the threshold contrast for a constant d' predicted in equation 113-14 has not been verified for human observers (Cohen, 1979). The effects of altered viewing conditions and training have yet to be investigated. Further discussion of this topic may be found in Chapter 115.

Three-dimensional aspects

The discussion of preceding sections dealt with the detection of a two-dimensional object in a single CT scan. In reality, however, the radiologic detection problem is three-dimensional in nature. The difficulties in detecting three-

*Personal communication, 1979.

dimensional objects in CT scans are often referred to as "partial volume" effects (Chapter 109). To illustrate the problem, consider the detection of a sphere of diameter d immersed in a uniform background of slightly lower density. If the sphere happened to lie completely within a single CT slice, its effective reconstruction density would be less than its actual density because of the partial volume effect. But if adjacent slices happened to split the sphere in half, its reconstructed density would be halved relative to the case just described. In the latter situation, the detection of the sphere is made more difficult by the large reduction in its reconstruction density, particularly if each slice is viewed independently. An improvement in detection could be attained by simply averaging the two slices, since the rms noise would be reduced by a factor of $\frac{1}{\sqrt{2}}$. This problem illustrates the desirability of a display system that allows full use of the three-dimensional information available in CT (Hanson, 1979a).

CONCLUSION

The detection limitations inherent in statistically limited computed tomographic (CT) images have been described through the application of signal detection theory. The detectability of large-area, low-contrast objects has been shown to be chiefly dependent on the low-frequency content of the noise power spectral density. For projection data containing uncorrelated noise, the resulting ramplike, low-frequency behavior of the noise power spectrum of the CT reconstruction may be conveniently characterized by the density of noise-equivalent quanta (NEQ) detected in the projection measurements. The NEQ for a given image can be determined either from a measurement of the noise power spectrum or from the noise granularity computed with an appropriate weighting function. The performance for the detection of large objects is as good in an efficient reconstruction (e.g., filtered back projection) as that based on the projection data. A measure of the efficiency of scanner dose utilization was presented that compares the average dose required by an ideal x-ray scanner to obtain the measured NEQ to that delivered by the actual scanner.

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Volume Five

Radiology of
the skull and brain

TECHNICAL ASPECTS OF
COMPUTED TOMOGRAPHY

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